## Outline

## Statistical Analysis of Corpus Data with R

You shall know a word by the company it keeps!
Collocation extraction with statistical association measures
— Part 2 -

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## Scaling up

- We know how to compute association scores ( $X^{2}$, Fisher, and $\log \theta$ ) for individual contingency tables now ...
... but we want to do it automatically for 24,000 bigrams in the Brown data set, or an even larger number word pairs
- Of course, you can write a loop (if you know C/Java):

```
> attach(Brown)
> result <- numeric(nrow(Brown))
> for (i in 1:nrow(Brown)) {
    if ((i %% 100) == 0) cat(i, " bigrams done\n")
    A <- rbind(c(O11[i],O12[i]), c(O21[i],O22[i]))
    result[i] <- chisq.test(A)$statistic
}
```

Scaling up: working with large data sets
Statistical association measures
Sorting and ranking data frames

The evaluation of association measures Precision/recall tables and graphs MWE evaluation in R

## Vectorising algorithms

- Standard iterative algorithms (loops, function calls) are excruciatingly slow in $R$
- $R$ is an interpreted language designed for interactive work and small scripts, not for implementing complex algorithms
- Large amounts of data can be processed efficiently with vector and matrix operations $\&$ vectorisation
- even computations involving millions of numbers are carried out instantaneously
- How do you store a vector of contingency tables?
as vectors $O_{11}, O_{12}, O_{21}, O_{22}$ in a data frame


## Vectorising algorithms

- High-level functions like chisq.test () and fisher.test () cannot be applied to vectors
- only accept a single contingency table
- or vectors of cross-classifying factors from which a contingency table is built automatically
- Need to implement association measures ourselves
- i.e. calculate a test statistic or effect-size estimate to be used as an association score
$\Rightarrow$ have to take a closer look at the statistical theory


## Adding marginals and expected frequencies in R

```
# first, keep R from performing integer arithmetic
> Brown <- transform(Brown,
    O11=as.numeric(011), 012=as.numeric(012),
    021=as.numeric(021), 022=as.numeric(022))
> Brown <- transform(Brown,
    R1=011+012, R2=021+022,
    C1=011+O21, C2=012+O22,
    N=O11+O12+O21+O22)
```

\# we could also have calculated them laboriously one by one:

```
Brown$R1 <- Brown$O11 + Brown$O12 # etc.
```

> Brown <- transform(Brown,
$\mathrm{E} 11=(\mathrm{R} 1 * \mathrm{C} 1) / \mathrm{N}, \mathrm{E} 12=(\mathrm{R} 1 * \mathrm{C} 2) / \mathrm{N}$,
$\mathrm{E} 21=(\mathrm{R} 2 * \mathrm{C} 1) / \mathrm{N}, \mathrm{E} 22=(\mathrm{R} 2 * \mathrm{C} 2) / \mathrm{N})$
\# now check that E11, ... E22 always add up to N!

Observed and expected frequencies

|  | $w_{2}$ | $\neg w_{2}$ |  |
| ---: | :--- | :--- | :--- |
| $w_{1}$ | $O_{11}$ | $O_{12}$ | $=R_{1}$ |
| $\neg w_{1}$ | $O_{21}$ | $O_{22}$ | $=R_{2}$ |
|  |  | $=N$ |  |


|  | $w_{2}$ | $\neg w_{2}$ |
| :---: | :---: | :---: |
| $w_{1}$ | $E_{11}=\frac{R_{1} C_{1}}{N}$ | $E_{12}=\frac{R_{1} C_{2}}{N}$ |
| $\neg w_{1}$ | $E_{21}=\frac{R_{2} C_{1}}{N}$ | $E_{22}=\frac{R_{2} C_{2}}{N}$ |

- $R_{1}, R_{2}$ are the row sums ( $R_{1}=$ marginal frequency $f_{1}$ )
- $C_{1}, C_{2}$ are the column sums $\left(C_{1}=\right.$ marginal frequency $\left.f_{2}\right)$
- $N$ is the sample size
- $E_{i j}$ are the expected frequencies under independence $H_{0}$


## Statistical association measures

Measures of significance

- Statistical association measures can be calculated from the observed, expected and marginal frequencies
- E.g. the chi-squared statistic $X^{2}$ is given by

$$
\text { chi-squared }=\sum_{i j} \frac{\left(O_{i j}-E_{i j}\right)^{2}}{E_{i j}}
$$

(you can check this in any statistics textbook)

- The chisq.test () function uses a different version with Yates' continuity correction applied:

$$
\text { chi-squared }_{\text {corr }}=\frac{N\left(\left|O_{11} O_{22}-O_{12} O_{21}\right|-N / 2\right)^{2}}{R_{1} R_{2} C_{1} C_{2}}
$$

## Statistical association measures

Measures of significance

- P-values for Fisher's exact test are rather tricky (and computationally expensive)
- Can use likelihood ratio test statistic $G^{2}$, which is less sensitive to small and skewed samples than $X^{2}$
(Dunning 1993, 1998; Evert 2004)
- $G^{2}$ uses same scale (asymptotic $\chi_{1}^{2}$ distribution) as $X^{2}$, but you will notice that scores are entirely different

$$
\text { log-likelihood }=2 \sum_{i j} O_{i j} \log \frac{O_{i j}}{E_{i j}}
$$

## Significance measures in R

Watch your numbers!

- $\log 0$ is undefined, so $G^{2}$ cannot be calculated if any of the observed frequencies $O_{i j}$ are zero
- Why are the expected frequencies $E_{i j}$ unproblematic?
- For these terms, we can substitute $0=0 \cdot \log 0$

```
> Brown <- transform(Brown,
    log}1=2 * 
        ifelse(O11>0, 011*log(011/E11), 0) +
        ifelse(O12>0, 012*log(012/E12), 0) +
        ifelse(O21>0, 021*log(O21/E21), 0) +
        ifelse(O22>0, O22*log(O22/E22), 0)
    ))
# ifelse() is a vectorised if-conditional
```


## Significance measures in R

```
# chi-squared statistic with Yates' correction
> Brown <- transform(Brown,
    chisq = N *
    (abs(011*O22 - O12*O21) - N/2)^2 /
    (R1 * R2 * C1 * C2)
)
```

\# Compare this to the output of chisq.test () for some bigrams.
\# What happens if you do not apply Yates' correction?
> Brown <- transform(Brown, $\log 1=2$ * $($
$011 * \log (011 / E 11)+012 * \log (012 / E 12)+$ $021 * \log (\mathrm{O} 21 / \mathrm{E} 21)+\mathrm{O} 22 * \log (\mathrm{O} 22 / \mathrm{E} 22)$
))
> summary (Brown\$logl) \# do you notice anything strange?

## Effect-size measures

- Direct implementation allows a wide variety of effect size measures to be calculated
- but only direct maximum-likelihood estimates, confidence intervals are too complex (and expensive)
- Mutual information and Dice coefficient give two different perspectives on collocativity:

$$
\mathrm{MI}=\log _{2} \frac{O_{11}}{E_{11}} \quad \text { Dice }=\frac{2 O_{11}}{R_{1}+C_{1}}
$$

- Modified log odds ratio is a reasonably good estimator:

$$
\text { odds-ratio }=\log \frac{\left(O_{11}+\frac{1}{2}\right)\left(O_{22}+\frac{1}{2}\right)}{\left(O_{12}+\frac{1}{2}\right)\left(O_{21}+\frac{1}{2}\right)}
$$

## Further reading

- There are many other association measures
- Pecina (2005) lists 57 different measures
- Evert, S. (to appear). Corpora and collocations.

In A. Lüdeling and M. Kytö (eds.), Corpus Linguistics. An International Handbook, article 57. Mouton de Gruyter, Berlin.

- explains characteristic properties of the measures
- contingency tables for textual and surface cooccurrences
- Evert, Stefan (2004). The Statistics of Word Cooccurrences: Word Pairs and Collocations.
Dissertation, Institut für maschinelle Sprachverarbeitung,
University of Stuttgart. Published in 2005, URN
urn:nbn:de:bsz:93-opus-23714.
- full sampling models and detailed mathematical analysis
- Online repository: www.collocations.de/AM
- with reference implementations in the UCS toolkit software
tall these sources use the notation introduced here


## How to use association scores

- Goal: use association scores to identify "true" collocations
- Strategy 1: select word pairs with score above threshold
- no theoretically motivated thresholds for effect size
- significance thresholds not meaningful for collocations (How many bigrams are significant with $p<.001$ ?)
- alternative: take $n=100,500,1000, \ldots$ highest-scoring word pairs $\stackrel{\circ}{ }$ n-best list (empirical threshold)
- Strategy 2: rank word pairs by association score
- reorder data frame by decreasing association scores
- word pairs at the top are "more collocational"
- corresponds to $n$-best lists of arbitrary sizes


## Implementiation of the effect-size measures

```
# Can you compute the association scores without peeking ahead?
> Brown <- transform(Brown,
    MI = log2(O11/E11),
    Dice = 2 * 011 / (R1 + C1),
    log.odds = log(
        ((011 + . 5) * (022 + .5)) /
        ((012 +.5)*(021 + .5))
    ))
```

\# check summary (Brown) : are there any more NA's?

## Rankings in R

```
> sum(Brown$chisq > qchisq(.999,df=1)) #p<.001
> sum(Brown$logl > qchisq(.999,df=1))
> Brown <- transform(Brown,
    r.logl = rank (-logl), # rank by decreasing score
    r.MI = rank(-MI, ties="min"), #see ?rank
    r.Dice = rank(-Dice, ties="min"))
> subset(Brown, r.logl <= 20, # 20-best list for log-likelihood
    c(word1,word2,011, logl,r.logl,r.MI,r.Dice))
```

\# Now do the same for MI and Dice. What are your observations?
\# How many anti-collocations are there among the 100 most \# collocational bigrams according to log-likelihood?

## Sorting data frames in R

$>x<-10$ * sample(10) \# 10, 20, ..., 100 in random order
$>$ sort ( x ) \# sorting a vector is easy (default: ascending)
$>$ sort(x, decreasing=TRUE)
\# But for sorting a data frame, we need an index vector that tell us \# in what order to rearrange the rows of the table.
> sort.idx <- order(x) \# also has decreasing option
$>$ sort.idx
$>x[$ sort.idx]

## Sorting data frames in R: practice time

Example solutions for practice questions

```
> paste(Brown.logl$word1, Brown.logl$word2) [1:100]
> paste(Brown$word1, Brown$word2)[sort.idx[1:100]]
# advanced code ahead: make your life easy with some R knowledge
> show.nbest <- function(myData,
    AM=c("chisq","logl","MI","Dice","O11"), n=20) {
        AM <- match.arg(AM) # allows unique abbreviations
        idx <- order(myData[[AM]], decreasing=TRUE)
        myData[idx[1:n], c("word1","word2","O11",AM)]
}
> show.nbest(Brown, "chi")
```

\# Can you construct a table that compares the measures side-by-side?

## Sorting data frames in R: practice time

\# try to sort bigram data set by log-likelihood measure
> sort.idx <- order(Brown\$logl, decreasing=TRUE)
> Brown.logl <- Brown[sort.idx, ]
> Brown. $\log 1[1: 20,1: 6]$
\# Now construct a simple character vector with the first 100 bigrams,
\# or show only relevant columns of the data frame for the first 100 rows.
\# Show the first 100 noun-noun bigrams (pos code N) and \# the first 100 adjective-noun bigrams (codes J and N ).
\# If you know some programming, can you write a function that \# displays the first $n$ bigrams for a selected association measure?

## Evaluation of association measures

- One way to achieve a better understanding of different association measures is to evaluate and compare their performance in multiword extraction tasks
- published studies include Daille (1994), Krenn (2000), Evert \& Krenn (2001, 2005), Pearce (2002) and Pecina (2005)
- "Standard" multiword extraction approach
- extract (syntactic) collocations from suitable text corpus
- rank according to score of selected association measure
- take $n$-best list as multiword candidates
- additional filtering, e.g. by frequency threshold
- candidates have to be validated manually by expert
- Evaluation based on manual validation
- expert marks candidates as true (TP) or false (FP) positive
- calculate precision of n-best list $=\# T P / n$
- if all word pairs are annotated, also calculate recall


## The PP-verb data set of Krenn (2000)

- Krenn (2000) used a data set of German PP-verb pairs to evaluate the performance of association measures
- goal: identification of lexicalised German PP-verb combinations such as zum Opfer fallen (fall victim to), ums Leben kommen (lose one's life), im Mittelpunkt stehen (be the centre of attention), etc.
- manual annotation distinguishes between support-verb constructions and figurative expressions (both are MWE)
- candidate data for original study extracted from 8 million word fragment of German Frankfurter Rundschau corpus
- PP-verb data set used in this session
- candidates extracted from full Frankfurter Rundschau corpus (40 million words, July 1992 - March 1993)
- more sophisticated syntactic analysis used
- frequency threshold $f \geq 30$ leaves 5102 candidates


## Table of $n$-best precision values

- Evaluation computes precision (and optionally) recall for various association measures and n -best lists

| n-best | logl | chisq | t-score | MI | Dice | odds | freq |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 100 | 42.0 | 24.0 | 38.0 | 19.0 | 21.0 | 17.0 | 27.0 |
| 200 | 37.5 | 23.5 | 35.0 | 16.5 | 19.5 | 14.0 | 26.5 |
| 500 | 30.4 | 24.6 | 30.2 | 18.0 | 16.4 | 19.6 | 23.0 |
| 1,000 | 27.1 | 23.9 | 28.1 | 21.6 | 14.9 | 24.4 | 19.2 |
| 1,500 | 25.3 | 25.0 | 24.8 | 24.3 | 13.2 | 25.3 | 18.0 |
| 2,000 | 23.4 | 23.4 | 21.9 | 23.1 | 12.6 | 23.3 | 16.3 |

- More intuitive presentation for arbitrary $n$-best lists in the form of precision graphs (or precision-recall graphs)


## Precision graphs



## Precision graphs: zooming in



The PP-verb data set

- krenn_pp_verb.tbl available from course homepage
- Data frame with 5102 rows and 14 columns:
- PP = prepositional phrase (lemmatised)
- verb = lexical verb (lemmatised)
- is.colloc = Boolean variable indicating TPs (= MWE)
- is.SVC, is.figur distinguish subtypes of MWE
- freq, MI, Dice, z.score, t.score, chisq, chisq.corr, log.like, Fisher = precomputed association scores (Do you recognise all association measures?)
- Our goal is to reproduce the table and plots shown on the previous slides (perhaps not all the bells and whistles)


## Precision-by-recall graphs



Precision tables: your turn!

```
> PPV <- read.delim("krenn_pp_verb.tbl")
> colnames(PPV)
> attach (PPV)
```

\# You should now be able to sort the data set and calculate \# precision for some association measures and n-best lists. \# (hint: sum () counts TRUE entries in Boolean vector)

## Precision tables

> idx.logl <- order(log.like, decreasing=TRUE)
$>\operatorname{sum}(i s . c o l l o c[i d x . \log l[1: 500]]) / 500 \quad \# n=500$
$>\operatorname{sum}(i s . c o l l o c[i d x . \log l[1: 1000]]) / 1000 \# n=1000$
\# use cumsum () to calculate precision for all n-best lists
> prec <- cumsum(is.colloc[idx.logl]) / (1:nrow(PPV))
$>\operatorname{prec}[c(100,200,500,1000,1500,2000)]$
\# first, generate sort index for each association measure
> idx.ll <- order(log.like, decreasing=TRUE)
> idx.chisq <- order(chisq, decreasing=TRUE)
> idx.t <- order(t.score, decreasing=TRUE)
> idx.MI <- order (MI, decreasing=TRUE)
> idx.Dice <- order(Dice, decreasing=TRUE)
> idx.f <- order(freq, decreasing=TRUE)

## Precision tables: an elegant solution

```
> n.list <- c(100,200,500,1000,1500,2000)
```

\# data frames of same height can be combined in this way
> prec.table <- cbind(
show. prec(PPV, "log.like", n.list), show. prec (PPV, "Fisher", n.list), show.prec (PPV, "chisq", n.list), show.prec(PPV, "chisq.corr", n.list), show.prec (PPV, "z.score", n.list), show.prec(PPV, "t.score", n.list), show.prec(PPV, "MI", n.list), show. prec (PPV, "Dice", n.list), show.prec (PPV, "freq", n.list)
)
> round (prec.table, 1) \# rounded values are more readable

Precision tables: an elegant solution

```
```

> show.prec <- function(myData, AM, n) {

```
```

> show.prec <- function(myData, AM, n) {
stopifnot(AM %in% colnames(myData)) \# safety first!
stopifnot(AM %in% colnames(myData)) \# safety first!
sort.idx <- order(myData[[AM]], decreasing=TRUE)
sort.idx <- order(myData[[AM]], decreasing=TRUE)
prec <- cumsum(myData$is.colloc[sort.idx]) /
    prec <- cumsum(myData$is.colloc[sort.idx]) /
(1:nrow(myData))
(1:nrow(myData))
result <- data.frame(100 * prec[n]) \#percentages
result <- data.frame(100 * prec[n]) \#percentages
rownames(result) <- n \# add nice row/column labels
rownames(result) <- n \# add nice row/column labels
colnames(result) <- AM
colnames(result) <- AM
result \# return single-column data frame with precision values
result \# return single-column data frame with precision values
}
}
> show.prec(PPV, "chisq", c(100,200,500,1000))

```
```

> show.prec(PPV, "chisq", c(100,200,500,1000))

```
```


## Precision graphs

## Precision graphs

\# second, calculate precision for all $n$-best lists
> n.vals <- 1:nrow (PPV)
> prec.ll <- cumsum(is.colloc[idx.ll]) *
100 / n.vals
> prec.chisq <- cumsum(is.colloc[idx.chisq]) * $100 / \mathrm{n} . \mathrm{vals}$
> prec.t <- cumsum(is.colloc[idx.t]) * $100 / \mathrm{n} . \mathrm{vals}$
> prec.MI <- cumsum(is.colloc[idx.MI]) * $100 / \mathrm{n} . \mathrm{vals}$
> prec.Dice <- cumsum(is.colloc[idx.Dice]) * 100 / n.vals
> prec.f <- cumsum(is.colloc[idx.f]) * $100 / \mathrm{n} . \mathrm{vals}$

## Precision graphs

\# increase font size, set plot margins (measured in lines of text)
$>\operatorname{par}(\operatorname{cex}=1.2, \operatorname{mar}=\mathrm{c}(4,4,1,1)+.1)$
\# third: plot as line, then add lines for further measures
$>$ plot(n.vals, prec.ll, type="l", ylim=c $(0,42), x a x s=" i ", ~ \# f i t x$-axis range tightly lwd=2, col="black", \# line width and colour xlab="n-best list", ylab="precision (\%)")
> lines(n.vals, prec.chisq, lwd=2, col="blue")
> lines(n.vals, prec.t, lwd=2, col="red")
> lines(n.vals, prec.MI, lwd=2, col="black", lty="dashed") \# line type: solid, dashed, dotted,...
> lines(n.vals, prec.Dice, lwd=2, col="blue", lty="dashed")
> lines(n.vals, prec.f, lwd=2, col="red", lty="dashed")

Precision graphs: playtime

- Add further decorations to plot (baseline text, arrows, ...)
- Write functions to simplify plot procedure
- you may want to explore type="n" plots
- Precision values highly erratic for $n<50 \Rightarrow$ don't show
- Graphs look smoother with thinning
- increment $n$ in steps of 5 or 10 (rather than 1)
- Calculate recall and create precision-by-recall graphs
all those bells, whistles and frills are implemented in the UCS toolkit (www.collocations.de/software.html)

